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# The impact of Covid-19 physical school closure on student performance in OECD countries: a meta-analysis

Di Pietro, G.

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#### **Contact information**

Name: Giorgio Di Pietro Address: Edificio Expo, c. Inca Garcilaso, 3, 41092 Seville (Spain) Email: giorgio.di-pietro@ec.europa.eu Tel.: (+34) 954488479

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# Abstract

In this report, we conduct a new meta-analysis of papers examining the impact of Covid-19 school related closure on student performance. While we focus only on OECD countries, the present meta-analysis includes, to the best of our knowledge, a larger number of studies (i.e., 55), effect sizes (i.e., 400) and countries (i.e., 21) than previous similar studies. Our results confirm that Covid-19 had, on average, an adverse effect on learning. While the size of the overall learning loss is estimated to be between 0.11 and 0.17 standard deviations of student achievement, learning losses are found to be smaller for pupils in OECD EU countries than for their peers in OECD non-EU countries. The periods of physical school closure were shorter in OECD EU countries than in OECD non-EU countries and this may provide a possible explanation for our finding. Additionally, our study shows that, overall, students seem to have fallen behind in their learning more in the later stages of the pandemic compared with the earlier stages. This finding is at variance with the outcome of earlier meta-analyses concluding that students did not lose any additional ground but failed to rebound. Our result is driven by the inclusion of recent studies showing that pandemic-related learning deficits have accelerated over time. Consequently, our findings suggest that particular attention should continue to be paid towards ensuring that students are able to catch up on what they have missed while schools have been closed.

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#### Authors

Giorgio Di Pietro, European Commission, Joint Research Centre, Seville

# 1 Introduction

Following the spread of Covid-19 in spring 2020, almost all educational institutions around the world were temporarily closed (UNESCO, 2020). Given the impossibility to conduct face to face education, students had to continue their learning online from home. However, the switch to remote education was sudden and many teachers and students were unprepared for this change. There is evidence showing that a lot of teachers did not possess the pedagogical and digitals skills required for online teaching (Newton, 2020). Similarly, a large number of students did not have the appropriate technological resources (e.g., laptop/tablet, stable internet connection) to follow virtual classes at home (Di Pietro et al., 2020; Blaskó et al., 2022). Additionally, some students had to share their computer with other family members and did not have a quiet space to study at home (Bolt, 2021). One should also consider that due to the pandemic children were confined to a long period of social isolation, which could have affected their well-being (European Commission/EACEA, 2022; Baarck et al., 2021). The increased risks of depression and anxiety are likely to have adversely impacted student performance given the close association between mental health and achievement (Wang et al., 2023).

Consistent with the above arguments, there is a large consensus, supported by empirical evidence, that Covid-19 school closure caused a learning loss among students. This is to say that students, on average, learned less during the pandemic relative to what they would have learned if the pandemic had not occurred. However, most of the existing evidence comes from single-country studies given the lack of relevant harmonised cross-country data, at least until recently (PIRLS 2021 microdata were realised in June 2023<sup>1</sup>). A strategy that can be employed to overcome this problem and hence deliver results of more general relevance is meta-analysis. Meta-analysis techniques combine and summarise the findings of individual national studies, thus producing outcomes with increased statistical power and less biases (Borenstein et al., 2021).

While there are already several systematic reviews (Patrinos et al., 2022; Moscoviz and Evans, 2022, Donnelly and Patrinos, 2021; Hammerstein et al., 2021; Zierer, 2021) and meta-analyses (Storey and Zhang, 2022; König and Frey, 2022; Di Pietro, 2023; Betthäuser et al., 2023) looking at the impact of the pandemic on student achievement, more research of this type is still needed. First, the literature in this area is expanding very quickly and evidence about the Covid-19 educational impact in certain countries has only been recently published for the first time. Second, given that most new studies consider results of assessments that took place one year or more after the first lockdown, updated meta-analyses allow us to investigate how learning losses have changed since the early stages of the pandemic. It is very important to understand if and to what extent students have been able to catch up with unfinished learning from the pandemic. Without effective remedial education, there could be long-term negative consequences for students and society. Students would enter the labour market with lower skills levels, and this would translate into lower employment opportunities and lower wages (Psacharopoulos et al., 2020). At macro-level, such a situation would have an adverse effect on economic growth (Hanushek and Woessmann, 2020).

We focus on OECD countries. The advantage of doing so is that these tend to be high income and upper-middle income countries with similar economic characteristics (Yilanci et al., 2019). On the other hand, however, our analysis does not capture the challenges faced by less developed countries whose students are especially likely to have suffered from a lack of devices, resources and internet connectivity, which might have prevented them from pursuing learning online. Despite considering only OECD countries, our meta-analysis does expand on previous meta-analyses and systematic reviews by using a larger number of effect sizes and studies<sup>2</sup>. Specifically, a total of 400 separate estimates from 55 studies are considered. Additionally, this meta-analysis covers a larger number of countries (i.e., 21) than previous meta-analyses<sup>3</sup>.

The remainder of this study is organised as follows. Section 2 describes the process of selecting studies and collecting data. It also discusses the empirical approach and the possibility of publication bias. Section 3 reports and discusses the main empirical results. Section 4 presents additional findings when analysing the educational

<sup>&</sup>lt;sup>1</sup> Kennedy and Strietholt (2023) use PIRLS 2021 data to examine the impact of the pandemic on student achievement.

<sup>&</sup>lt;sup>2</sup> Previous meta-analyses include König and Frey (2022) who extracted 109 effect sizes nested in 18 studies, Storey and Zhang (2021) who synthetized 79 effect sizes from 10 studies, Betthäuser et al. (2023) who considered 291 effect sizes from 42 studies, and Di Pietro (2023) who considered 239 effect sizes from 39 studies. The reviews by Patrinos et al. (2022), Moscoviz and Evans (2022), Donnelly and Patrinos (2021), Hammerstein et al. (2021) and Zierer (2021) summarised the results of 35 studies, 29 studies, 8 studies, 11 studies and 9 studies, respectively.

<sup>&</sup>lt;sup>3</sup> Both König and Frey (2022) and Storey and Zhang (2021) considered 5 countries. While the meta-analysis by Betthäuser et al. (2023) covered 15 countries, Di Pietro (2023) collected relevant information from studies conducted in 19 different countries.

impact of Covid-19 broken down by gender and socio-economic status. Section 5 looks at remedial policies designed to alleviate the learning losses due to the pandemic. Section 6 concludes.

# 2 Method

To carry out this meta-analysis, we followed the Preferred Reporting Items for Systematic Reviews and Metaanalysis (PRISMA) (Moher et al., 2009).

# 2.1 Inclusion criteria

With the purpose of this study in mind, a set of inclusion criteria was defined. They guided the selection of the studies included in this meta-analysis. Specifically, the following four inclusion criteria were used:

• the study should quantitatively examine the effect of Covid-19 on student achievement in primary, secondary or tertiary education in an OECD country. This means that the data used in this study were collected before and during the pandemic (or only during the pandemic if, when schools were closed, some students were still receiving in-person teaching thereby simulating pre-pandemic conditions), therefore clearly distinguishing between a control and a treated group, respectively.

• the study should use objective indicators (e.g., test scores) to measure student achievement.

• the study should be based on real data.

• the study should report data on an effect size (or sufficient information to compute it) and its standard error (or *t*-statistic, *p*-value, or sufficient information to calculate it).

# 2.2 Search and screening procedures

In order to ensure a comprehensive and multidisciplinary search, we consulted five electronic databases (i.e., Google Scholar, EconLit, ScienceDirect, Education Resources Information Center, and PubMed). We employed the following keywords: "Covid-19 (OR coronavirus OR pandemic OR Cov) AND student (OR academic OR scholastic) performance (OR achievement OR learning OR outcome) OR test score".

This search, which ended on 6<sup>th</sup> March 2023, yielded thousands of results. After removing duplicates, we screened the title and abstract of each remaining item. The large majority of these studies were excluded because they use qualitative approaches (e.g., interviews), report teachers'/parents' views about the educational impact of Covid-19 (e.g., Kim et al., 2023; Lupas et al., 2021), focus on non-OECD countries (e.g., El Said, 2021; Fen et al., 2021), or provide a theoretical discussion about how the pandemic is likely to affect education (e.g., Di Pietro et al., 2020). Similarly, studies containing predictions and/or projections were also removed (Kuhfeld et al., 2020a). After this initial screening, the content of the remaining 196 studies was carefully examined, and only those fulfilling all the inclusion criteria were considered. In this phase, we excluded studies that, although attempting to understand how the pandemic impacted student learning, employ a different outcome measure (e.g., dropout rate) than the one considered in this meta-analysis (e.g., Tsolou et al., 2021). In the same vein, we removed studies using student self-reported outcome measures as well as those examining the educational impact of Covid-19 on specific subgroups of students (e.g., Agostinelli et al., 2022). Additionally, in order to ensure that key sources were not missed, we consulted reference lists included in previous metaanalyses/systematic reviews and contacted colleagues and researchers who have an interest in this area. 3 more relevant studies were identified in this way. A total of 55 studies was included in this study. Figure 1 summarizes the literature search and the screening procedure.

#### Figure 1: Flowchart illustrating the review selection process



# 2.3 Study coding

All the studies included in this meta-analysis were read in-depth, and relevant information and findings were extracted.

In line with the current best practice in meta-analysis (Polák, 2019), we use all relevant estimates included in the selected studies. As argued by Cheung (2019), not doing so results in missed opportunities to take advantage of all the available data to answer the research question/s under investigation. However, a fundamental issue with this approach lies in the dependence between multiple estimates from the same study given that effect sizes are assumed to be independent in meta-analysis (Cheung and Vijayakumar, 2016). As discussed later in the report (see Section 2.4), the Robust Variance Estimation (RVE) approach is used to account for within-study dependence.

# 2.3.1 Effect size calculation

In order to be able to aggregate the various impact estimates reported in the selected studies, one needs to convert them into a common metric. Consistent with previous relevant systematic reviews and meta-analyses, the Cohen's *d* as a scale-free effect size measure is used. Cohen's *d* refers to standardised mean differences and is calculated by dividing the mean difference in student performance between pre-Covid and Covid periods by the pooled standard deviation. While in some cases the Cohen's *d* was retrieved from the studies, in others it was calculated using information directly available from them. If not reported, the Cohen's *d* standard error was computed using the formula given in Cooper and Hedges (1994). In case no information on sample sizes were available from the studies but exact *p*-values were instead reported, the formula provided by Higgins and Green (2011) was employed to obtain standard errors. In some instances, we also used information on effect sizes contained in the electronic supplements of the meta-analysis articles by König and Frey (2022), Di Pietro (2023), and Betthäuser et al. (2023). For instance, this was the case when a study does not report Cohen's *d* but this information has been already collected by König and Frey who have contacted the relevant author/s.

# 2.3.2 Moderator variables

Six variables were coded as possible moderators:

#### a) Subject area

Following the approach of Di Pietro (2023), three different subject areas are selected: humanities, math/science and a mixed category. Several studies (e.g., Imberman and Strunk, 2023) suggest that the pandemic had a detrimental effect especially on math achievement. During the lockdown, parents might have found more difficult to support their children with math learning compared to language and literacy.

#### b) Geographical area

We distinguish between EU OECD countries (Italy, Germany, Netherlands, Finland, Sweden, Spain, Belgium, Hungary, Denmark, Poland, Greece, Czech Republic and Slovenia) and non-EU OECD countries (UK, US, Mexico, Turkey, Australia, Colombia, Norway and Switzerland).

#### c) Level of education

We differentiate between three levels of education: primary, secondary and tertiary. It is plausible that the level of education may serve as moderator as the pandemic could have had a detrimental impact especially on the achievement of children of younger age given that remote learning is particularly challenging for this group of students (Hinton, 2020).

# d) Timing of student assessment during Covid-19

In order to investigate if and to what extent students recovered from the learning loss caused by virtual learning, it is important to look at how the impact of Covid-19 on student achievement varied across the different stages of the pandemic. Specifically, we distinguish between assessments that have taken place in the early stages of Covid-19 (i.e., 2020) and those that occurred in later stages (i.e., 2021 or after).

# e) Type of publication

We distinguish between peer reviewed journal articles and other studies. Publication type is a common moderator variable in meta-analysis. Peer reviewed journal articles are expected to be of higher scientific rigour and are less likely to contain typos or errors in the reported results.

# f) Publication year

Publication year is another typical moderator variable in meta-analysis. In line with the approach followed by several meta-analyses (e.g., Di Pietro, 2022; Havranek and Irsova, 2017), the year of the first appearance of a draft of the study in Google Scholar is used. This measure is preferred to the nominal publication year because of increasing publication lags (Zigraiova et al., 2020).

# 2.4 Estimators and models

Two approaches frequently used in the meta-analysis literature are: 1) the Fixed Effects (FE) model, and 2) the Random Effects (RE) model. They rely on different assumptions. The FE model assumes that there is one true effect size common to all studies and that all differences in the observed effects can be attributed to within-study sampling error. By contrast, the RE model states that the effect size may vary between studies not only due to within-study sampling error, but also because there is heterogeneity in true effects between studies. Such additional variability is typically modelled employing a between-study variance parameter. Considering the differing characteristics of the studies included in our sample, it seems unreasonable to assume that there is a common true effect that every study shares. Hence, it is anticipated that the RE model would be more suitable.

In our empirical analysis, we use the RE model to: a) compute estimates of the overall effect size, and b) run moderator analyses to explore potential sources of heterogeneity of results.

Following the suggestion of Veroniki et al. (2019), we first compute pooled effect sizes using several estimators (i.e., Restricted Maximum Likelihood (RML), DerSimonian and Laird (DL), Paule-Mandel (PM), and Sidik-Jonkman (SJ) that adopt different methods to estimate the between-study variance component. Furthermore, we apply the so-called Knapp-Hartung adjustment (Knapp and Hartung, 2003) to obtain more reliable estimates of the standard errors (and thus the confidence intervals) of the pooled effect sizes. The advantages related to the Knapp-Hartung method are outlined in several studies (see, for instance, Langan et al., 2019).

Additionally, to test the robustness of the above results, in line with the approach of several papers (e.g., Kaiser and Menkhoff, 2020), an overall effect size is also computed using the RVE method, which accounts for the possibility that multiple effect sizes from the same study are non-independent. The benefits of this method are that there is no need to drop any effect size (to ensure their statistical independency) and no information is required about the intercorrelation between effect sizes within studies.

Next, in an attempt to identify factors that may explain why the extent of the Covid-19 learning loss is found to vary across the included studies, moderator analyses are performed. Specifically, in order to test for the significance of categorial moderator variables, we compute Q-between (Q<sub>B</sub>), an index of the variability between the group means. Q<sub>B</sub> has a  $\chi^2$  distribution with L-1 degrees of freedom (where L is the number of groups) and its null hypothesis is the homogeneity of the groups. Q<sub>B</sub> statistics assess the amount of variance that can be attributed to between-group differences (Sirin and Sin, 2023). A significant Q<sub>B</sub> would suggest that the mean effect sizes across categories differ by more than sampling error. Analyses for moderators (i.e., level of education, subject area, publication year, publication type, timing of student assessment during Covid-19 and geographical area) are basically analogous to subgroup analyses of variance.

# 2.5 Sample characteristics

The dataset employed for the meta-analysis comprises a total of 400 different impact estimates extracted from 55 separate studies. Each study contained in the dataset includes a number of estimates that vary from 1 to 32. Several reasons explain why most studies (i.e., 87.3%) reported multiple estimates. Many studies (e.g., Maldonado and De Witte, 2022; Gambi and De Witte, 2021; Borgonovi and Ferrara, 2022; Bielinski et al., 2021; Vegas, 2022) estimated the effect of Covid-19 on student performance in several subjects. Similarly, a large number of studies (e.g., Contini et al., 2021; Gore et al., 2021; Domingue et al., 2021) examined the impact of the pandemic on the achievement of students of different levels of education or even of students of different grades within the same level of education. For instance, Meeter (2021) analysed how Covid-19 affected the math performance of primary school children of grades 2-6. Some studies also provided different estimates showing both the short and long-term effects of Covid-19 on student achievement. For example, Kuhfeld et al. (2022) looked at changes in student test scores in autumn 2020 and autumn 2021 relative to autumn 2019.

Table 1 presents the studies included in the dataset. Studies are listed alphabetically. For each study, we report information on the author(s), year of publication<sup>4</sup>, country examined, number of the effect sizes collected and

<sup>&</sup>lt;sup>4</sup> In this table, we report the actual year of publication of the latest version of the study (for journal articles this is the year when they are assigned a volume and issue number) rather than the year of the first appearance of a draft of the study in Google Scholar.

their mean value. The studies cover a total of 21 countries. The largest source countries are the US (129 estimates), the UK (58 estimates), Belgium (48 estimates) and Germany (41 estimates).

Study (Author(s) and year of	Country	Number of effect sizes	Mean
publication)		collected	effect size
Arenas and Gortazar (2022)	Spain	4	-0.04
Bazoli et al. (2022)	Italy	6	-0.16
Bielinski et al. (2021)	US	16	-0.13
Birkelund and Karlson (2021)	Denmark	5	0.05
Blainey and Hannay (2021a)	UK	12	-0.06
Blainey and Hannay (2021b)	UK	12	-0.19
Blainey and Hannay (2021c)	UK	12	-0.12
Borgonovi and Ferrara (2022)	Italy	4	-0.04
Contini et al. (2021)	Italy	2	-0.21
Coskun and Kara (2022)	Turkey	1	-0.34
De Paola et al. (2022)	Italy	1	-0.11
Department for Education (2021a)	UK	6	-0.11
Department for Education (2021b)	UK	2	-0.17
Depping et al. (2021)	Germany	32	-0.01
Domingue et al. (2021)	US	4	-0.03
Engzell et al. (2021)	Netherlands	4	-0.08
Fälth et al. (2021)	Sweden	18	0.09
Förster et al. (2023)	Germany	2	0.06
Gambi and De Witte (2021)	Belgium	22	-0.13
Gambi and De Witte (2023)	Belgium	15	-0.21
Gillitzer and Prasad (2022)	Australia	9	-0.02
GL Assessment (2021)	LIK	4	-0.22
Gore et al. (2021)	Australia	4	0.04
Haelermans et al. (2021)	Netherlands	3	-0.12
Hevia et al. (2022)	Mexico	2	-0.54
lakubowski et al. (2022)	Poland	2	-0.30
Kofoed et al. (2021)		5	-0.22
Koran and Lavertu (2021a)	115	1	-0.23
Kogan and Lavertu (2021b)	115	11	-0.23
Korbel and Prokon (2021)	Czech Benublic	2	-0.08
Korakowski et al. (2021)		12	-0.13
Kubfeld et al. $(2020h)$	115	12	-0.15
Kuhled et al. $(2020b)$	115	24	-0.10
$\frac{1}{2022}$	Grooco	24	-0.12
Larkkapon et al. (2022)	Einland	11	-0.20
Lerkkallell et al. $(2023)$		11	-0.12
Ludowig at al. $(2021)$	Gormany	2	-0.14
Ludewig et al. (2022)	Bolgium	11	-0.17
Mantor (2021)	Deigiuiii	11	-0.16
Melef (2021)	Nethenanus	10	0.15
Molinar and Hermann (2023)	Hungary	18	-0.11
Moliner and Alegre (2022)	Spain		-2.34
wonner et al. (2022)	Spain		-0.95
Driov et al. (2021)	05	2	-0.12
Pier et al. (2021)	US	25	-0.14
Reiyea et al. (2022)	US	3	-0.36
коse et al. (2021)	UK	6	-0.1/
Schult et al. (2021)	Germany	3	-0.06
Schuurman et al. (2023)	Netherlands	2	-0.08

 Table 1: Sources for meta-analysis

Skar et al. (2021)	Norway	2	-0.48
Spitzer and Musslick (2021)	Germany	2	0.15
Tomasik et al. (2021)	Switzerland	2	-0.07
van Der Velde (2021)	Netherlands	1	0.25
Vegas (2022)	Colombia	6	-0.27
Weidmann et al. (2021)	UK	4	-0.01
Žnidaršič et al. (2022)	Slovenia	1	0.11

Table 2 provides descriptive statistics for our moderator variables. While Column 1 reports simple averages (and standard deviations), Column 2 shows averages (and standard deviations) weighted by the inverse of the number of estimates reported in each study.

#### Table 2: Descriptive statistics

	Unweighted	Weighted (by the inverse of the
	Mean (Standard deviation)	number of estimates reported in
		each study)
Variable name		Mean (Standard deviation)
	(1)	(2)
Effect size (Cohen's d)	-0.113 (-0.010)	-0.174 (0.018)
Year of publication		
2020	0.093 (0.015)	0.073 (0.013)
2021	0.606 (0.024)	0.545 (0.025)
2022	0.178 (0.019)	0.291 (0.023)
2023	0.123 (0.106)	0.091 (0.014)
Subject area		
Math/Science	0.444 (0.025)	0.421 (0.025)
Humanities	0.528 (0.025)	0.506 (0.025)
Mix	0.028 (0.008)	0.073 (0.013)
Level of education		
Primary	0.670 (0.024)	0.605 (0.024)
Secondary	0.308 (0.023)	0.322 (0.023)
Tertiary	0.022 (0.007)	0.073 (0.013)
Timing of student assessment during	Covid-19	
2020	0.488 (0.025)	0.530 (0.025)
2021 or later	0.512 (0.025)	0.470 (0.025)
Geographical area		
OECD EU Member States	0.423 (0.025)	0.491 (0.025)
OECD non-EU Member States	0.577 (0.025)	0.509 (0.025)
Type of publication		
Peer reviewed journal articles	0.203 (0.020)	0.345 (0.024)
Other studies	0.797 (0.020)	0.655 (0.024)

# 2.6 Publication bias

Publication bias is a serious issue in meta-analysis, which can undermine the validity and generalization of the conclusions (Lin and Chu, 2018). It occurs because there is a tendency for researchers to submit or scientific journals to publish articles producing statistically significant results as opposed to null or uncertain results (Ji et al., 2011).

We use the Doi plot to visually assess publication bias. It employs a folded normal quantile (Z-score) versus effect plot. The main advantages of the Doi plot are that it aids visualization of asymmetry (in the absence of publication bias there is no asymmetry), and it quantifies the asymmetry through the Luis-Furuya-Kanamori (LFK) index. LFK

index values within ±1 suggest no asymmetry, LFK index values exceeding ±1 but within ±2 indicate minor asymmetry, while LFK index values exceeding ±2 denote major asymmetry (Furuya-Kanamori et al., 2018). As shown in Figure 2, the Doi plot shows no asymmetry (LFK index=0), indicating that no publication bias is detected. Our findings are consistent with those in previous relevant meta-analyses. König and Frey (2022), Di Pietro (2023), and Betthäuser et al. (2023) conclude that the presence of publication bias is unlikely.



Figure 2: Doi plot

# 3 Main results

We present results in two steps. First, we analyse overall effect sizes (Section 3.1). Second, we run moderator analyses to investigate sources of heterogeneity in effect sizes (Section 3.2).

# 3.1 Overall effect sizes

Table 3 shows summary statistics of the overall effect sizes employing the RE model. Column 1 of Table 3 reports the results using the RML estimator. The point estimate of the overall effect size turns out to be -0.108 (SE=0.009, *p*-value=0.000, 95% CI [-0.125, -0.091]). As shown in Columns 2, 3, and 4 of Table 3, results do not significantly change when other RE estimators are employed. More precisely, the point estimate of the overall effect size ranges between -0.105 (using DL) and -0.109 (using both PM and SJ), and it is always highly statistically significant. Similarly, regardless of the RE estimator used, the relevant point estimate continues to be highly statistically significant even when the Hartung-Knapp method is employed to compute the standard error of the pooled effect size. Additionally, both Cochran's Q test and the I<sup>2</sup> statistic indicate considerable heterogeneity across the results of the selected studies, thus confirming the appropriateness of the RE model. I<sup>2</sup>, which is an index typically used to quantify the degree of heterogeneity in a meta-analysis (Higgins and Thompson, 2002), has consistently a value of 100% across the RE estimators. This suggests that all the variability in the effect-size estimates is due to heterogeneity as opposed to sampling error.

_ · · · · · · · · · · · · · · · · · · ·				
	Restricted	Sidik-	DerSimonian	Paule-Mandel
	Maximum	Jonkman	and Laird	(PM)
	Likelihood	(SJ)	(DL)	
	(RML)			
	(1)	(2)	(3)	(4)
Effect size	-0.108	-0.109	-0.105	-0.109
(Standard Error)	(0.009)	(0.009)	(0.006)	(0.009)
[p-value]	[0.000]	[0.000]	[0.000]	[0.000]
95% Confidence Interval	-0.125, -0.091	-0.127, -0.091	-0.116, -0.095	-0.127, -0.091
95% Hartung–Knapp Confidence Interval	-0.126, -0.090	-0.127, -0.091	-0.122, -0.088	-0.127, -0.091
Q-statistic	430000000	430000000	430000000	430000000
[p-value]	[0.000]	[0.000]	[0.000]	[0.000]
l <sup>2</sup> (%)	100	100	100	100

 Table 3: Summary statistics of the overall effect sizes- Random Effects model

Next, in order to address within-study dependence in effect sizes, the RVE approach is employed. We fit an intercept-only RE RVE model to our set of effect sizes to compute the overall summary effect. In this model, the intercept can be interpreted as the precision-weighted mean effect size adjusted for effect-size dependence (Friese et al., 2017). The RVE RE mean effect size turns out to be  $-0.1735^5$  (SE=0.055, *p*-value=0.0005, 95% CI [-0.2676, -0.0795]), which is larger than our previous estimates. It is also important to note that in this model the small-sample corrected degrees of freedom is greater than 4 (i.e., 54), suggesting that the *p*-value for the associated *t*-test accurately reflects the type I error (Tanner-Smith et al., 2016).

One should observe that our findings from the RVE analysis are broadly consistent with those from previous meta-analyses. Storey and Zhang (2021) concluded that due to Covid-19 students lost, on average, 0.15 standard deviations of learning, König and Frey (2022) found average losses of 0.175 standard deviations, Betthäuser et al. (2023) estimated average losses at 0.14 standard deviations, and Di Pietro (2023) calculated average losses of

<sup>&</sup>lt;sup>5</sup> The *robumeta* command in Stata is employed. An intercept-only model is run where the estimate of the meta regression constant is equal to the unconditional mean effect size across studies. With this command, it is possible to specify a value for *rho*, the expected correlation among dependent effects. Following Tanner-Smith and Tipton (2013), we use different values of *rho* ranging from 0 to 1 in intervals of 0.1 in an attempt to check the consistency of results. All models yield the same outcome regardless of the specified value of *rho*.

around 0.19 standard deviations<sup>6</sup>. Two considerations help put these results into perspective. First, one may notice that the delayed learning suffered by students as a result of Covid-19 school closure is roughly comparable to that experienced by their peers after major natural disasters. For instance, Sacerdote (2012) found that in the spring of 2006 students who were displaced by Katrina and Rita hurricanes saw their test scores fall by between 0.07 and 0.2 standard deviations. A similar result, though of a smaller magnitude, is obtained by Thamtanajit (2020). He showed that in Thailand floods reduced student test scores by between 0.03 and 0.11 standard deviations, depending on the subject and educational level. Second, following Hanushek and Woessmann (2020), a learning deficit of about 0.186 standard deviations can be considered to be equivalent to the loss of just over half of a school year<sup>7</sup>.

# 3.2 Moderator analyses

The next step in the meta-analysis is to attempt to investigate factors that may explain the observed variation in effect sizes given that, as shown above,  $I^2$  is equal to 100%. In fact, when  $I^2$  is greater than 75% heterogeneity is high and this justifies moderator analyses (Higgins et al., 2003).

Table 4 reports the results of the moderator analyses<sup>8</sup>. It is interesting to observe that 95% CI of each subset of all moderator variables is negative and exclusive of zero, indicating that the pandemic, on average, had a consistent, statistically significant, detrimental effect on student achievement. Additionally, one should note that the categories of all moderator variables, with the exception of "Tertiary", are highly heterogeneous ( $l^2$ >99%). This means that practically all between-study heterogeneity is still unexplained.

# a) Subject areas

There are a total of 400 effect sizes that provide information on the subject being assessed. Of the 400 effect sizes, 211 represent humanities, 178 represent math/science and 11 represent a mix of these two subjects.

The average effect size for math/science (d = -0.128) is found to be larger than the corresponding ones for humanities (d = -0.093) and the mix category (d = -0.079). While on the one hand this result is consistent with those obtained in the meta-analyses by Di Pietro (2023) and Betthäuser et al. (2023), on the other, it differs because our average effect sizes are not statistically different. The small number of observations in the mix category may have contributed to the overlapping of the confidence intervals and statistical insignificance.

# b) Geographical area

There are a total of 400 effect sizes that give information on the geographical area where the relevant sample was collected. Out of the 400 effect sizes, 231 refer to samples drawn from OECD countries that are not part of the EU, while the remaining 169 represent samples drawn from OECD-EU Member States.

Our analysis reveals statistically significant moderating effects of geographical area. Specifically, Covid-19 caused a larger average learning loss in OECD non-EU countries (d = -0.139, Cl95 = -0.154, -0.123) than in OECD-EU Member States (d = -0.069, Cl95 = -0.104, -0.034).

A possible explanation for the smaller learning losses experienced by students in OECD EU countries relative to their peers in OECD non-EU countries may lie in differences in their exposure to physical school closure. As expected, Kennedy and Strietholt (2023) show that there is a negative correlation between physical school closure and student achievement in reading. Using data gathered by UNESCO's Institute of Statistics on the duration of Covid-19 physical school closure in many countries around the world, we computed the average number of days schools were closed (partially or totally) in our two groups of countries. While in the OECD EU countries included in our sample schools were physically closed due to Covid-19, on average, for about 245 days, the corresponding

<sup>&</sup>lt;sup>6</sup> Relevant systematic reviews have also found similar learning deficits. Donnelly and Patrinos (2022) found average delays of 0.13 standard deviations, Zierer (2021) estimated average losses at 0.14 standard deviations, and Hammerstein et al. (2021) reported average deficits of 0.10 standard deviations. Examining only EU countries, De Witte and François (2023) computed an average learning deficit of around 0.11 standard deviations.

<sup>&</sup>lt;sup>7</sup> They found that the loss of one third of a school year of learning is equivalent to approximately 11% of a standard deviation of lost test results. This finding is broadly consistent with that obtained by Hill et al. (2008) who conclude that a value of Cohen's *d* of 0.4 (with a margin of error of ± 0.06) corresponds to the average annual reading achievement gains in fourth grade.

<sup>&</sup>lt;sup>8</sup> The RML method is used.

value for OECD non-EU countries is 344 days. Interestingly, this gap widens if a weighted average (taking into account the number of effect sizes for each country within the two groups) is calculated (i.e., 422 days for OECD non-EU countries vs 228 days for OECD EU countries).

Table 4: Results for moderator variab	les
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	1	1	r		-
Moderator variables	QB	K	<i>d</i> +	95% CI	l <sup>2</sup>
Subject area	4.46				
1. Humanities		211	-0.093	-0.112 -0.074	100.00
2. Math/Science		178	-0.128	-0.160 -0.097	100.00
3. Mix	]	11	-0.079	-0.124 -0.034	99.84
Geographical area	12.75*				
1. OECD EU Member	1	169	-0.069	-0.104 -0.034	100.00
States					
2. OECD non-EU Member	1	231	-0.139	-0.154 -0.123	100.00
States					
Level of education	1.87				
1. Primary	1	268	-0.104	-0.123 -0.084	100.00
2. Secondary	1	123	-0.117	-0.153 -0.081	100.00
3. Tertiary	1	9	-0.144	-0.203 -0.086	45.24
Type of publication	0.38				
1. Peer reviewed journal	1	81	-0.129	-0.203 -0.056	100.00
articles					
2. Other studies	]	319	-0.105	-0.119 -0.092	100.00
Year of publication					
2020	9.95*	37	-0.120	-0.153 -0.087	100.00
2021	]	243	-0.087	-0.107 -0.067	100.00
2022	]	71	-0.156	-0.218 -0.093	99.99
2023	1	49	-0.151	-0.200 -0.102	94.91
Timing of student	13.72*				
assessment during Covid-					
19					
2020		195	-0.076	-0.099 -0.053	100.00
2021 or later		205	-0.139	-0.163 -0.115	99.94

Note:  $Q_B$  refers to the moderator contrasts, K = Number of effect sizes,  $d_*$  = average value of effect size, 95%CI = 95% Confidence Interval for the average value of effect size,  $l^2$  refers to the heterogeneity of each group; \* p < 0.05.

# c) Level of education

There are a total of 400 effect sizes that provide information on the education level of assessed students. Out of the 400 effect sizes, 268 represent primary students, 123 represent secondary students and 9 represent tertiary students.

 $Q_B$  turns out not to be significant in level of education (Q<sub>B</sub>=1.87, p >.05), meaning that the impact of the pandemic on student performance is not different across primary, secondary and tertiary students. This finding is in line with that of Di Pietro (2023). Betthäuser et al. (2023) also find that the impact of Covid-19 on achievement does not statistically differ between primary and secondary students. However, our result needs to be interpreted with caution given the small number of effect sizes related to tertiary students.

#### d) Timing of student assessment during Covid-19

There are a total of 400 effect sizes that provide information on the timing of student assessment during Covid-19. Out of the 400 effect sizes, 195 refer to assessments that took place in 2020 while the remaining 205 are from assessments that occurred in 2021 or later.

Our analysis indicates that the timing of student assessment during Covid-19 moderates the effect of the pandemic on student performance ( $Q_B$ =13.72, p < .05). The average loss in student achievement is smaller in assessments that took place in the early stages of the pandemic (i.e., 2020- *d*=-0.076) compared to assessments that occurred several months or more than one year after the first lockdown (i.e., 2021 or later- *d*=-0.139).

Therefore, our results suggest that, on average, students have fallen further behind since returning to the classroom following the pandemic. Many students might have been unable to grasp more complex academic concepts because they missed out on key foundational knowledge as a result of school closure during Covid-19.

The above outcome is consistent with that of Gambi and De Witte (2023) who find that in the Flemish Community of Belgium learning deficits seem to have accelerated over time. Increased teacher shortage caused by the pandemic could have contributed to this situation. Using data from 45 countries, most of which are high-income OECD members, Özdemin et al. (2022) estimate that the impact of the Covid-19 pandemic could erase decades-long gains in adult skills for the affected cohorts unless remedial measures are quickly adopted.

#### e) Type of publication

Type of publication is not found to have any significant relation with effect sizes for the impact of Covid-19 on student performance ( $Q_B = 0.38$ , p > .05). Whether the relevant analysis is included in peer reviewed journal articles (d = -0.129, Cl95 = -0.203, -0.056) or in other studies (d = -0.105, Cl95 = -0.119, -0.092) do not seem to affect the effect size. In fact, the confidence intervals for "Peer reviewed journal articles" and "Other studies" strongly overlap.

#### f) Publication year

Our analysis suggests that publication year is a significant moderator of the relationship between Covid-19 and student performance ( $Q_B = 9.95$ , p < .05). Studies published in 2022 (d = -0.156, Cl95 = -0.218, -0.093) or 2023 (d = -0.151, Cl95 = -0.200, -0.102) find larger average learning deficits than studies published in 2020 (d = -0.120, Cl95 = -0.153, -0.087) or 2021 (d = -0.087, Cl95 = -0.107, -0.067). One possible explanation for this finding is that more recent studies are more likely to examine results of assessments that took place one year or more after the first lockdown, and previous results indicate that learning losses accumulate over time.

# 4 Additional findings

Next, in an attempt to examine whether the impact of Covid-19 on student achievement does vary by gender and socio-economic background, we searched for relevant effect sizes in the 55 studies included in our sample. We found that separate gender estimates are reported in 15 studies while 27 studies look at the differential impact by socio-economic group. Following the approach of Card et al. (2010), we adopt a 'sign and significance' approach. Therefore, we classify gender effect sizes into three categories: statistically significant larger learning losses for boys, gender learning losses insignificantly different from zero, and statistically significant larger learning learning losses for girls. Similarly, effect sizes by socio-economic status are also divided into three categories: statistically significant larger learning losses for students from lower socio-economic status, statistically insignificant learning losses for students from lower socio-economic status.

The simple classification 'sign and significance' enables us to draw comparison across studies that use different definitions of socio-economic status<sup>9</sup>. Additionally, in several studies (e.g., Maldonado and De Witte, 2022) exact values for relevant effect sizes and the corresponding standard errors are not provided, but results are presented through graphs. Statistical significance is, however, indicated in the main text.

We collected 34 separate effect sizes for gender and 133 independent effect sizes for lower socio-economic status. Results are reported in Table 5. As for gender, the majority of effect sizes (i.e., 58.8%) support the hypothesis that the effect of the pandemic on student achievement did not statistically vary by gender. This finding is in line with the conclusion of De Witte and François (2023). They observe mixed evidence about the gender effects of Covid-19 on student performance. Table 5 also shows that most effect sizes (i.e., 54.1%) are associated with statistically significant larger learning losses for students from lower socio-economic status. While such an outcome is in line with previous studies (e.g., Betthäuser et al., 2023), one should note that it does not appear to be supported by an overwhelming body of evidence as the proportion of statistically insignificant effects (i.e., 37.6%) is quite high.

Gender	Statistically significant larger	Gender learning losses	Statistically significant larger
	learning losses for boys	statistically insignificant	learning losses for girls
	14.7%	58.8%	26.5%
Socio-	Statistically significant larger	Statistically insignificant	Statistically significant smaller
economic	learning losses for students	learning losses by socio-	learning losses for students
status	from lower socio-economic	economic status	from lower socio-economic
	status		status
	54.1%	37.6%	8.3%

Table 5: Relation between the educational impact of Covid-19 and gender/socio-economic status

<sup>&</sup>lt;sup>9</sup> A range of different definitions of socio-economic status is used. Some of them are based on parental income (e.g., Birkelund and Karlson, 2021), parental education (e.g., Engzell et al., 2021), under-represented minorities (e.g., Orlov et al., 2021), eligibility for a free meal (e.g., Rose et al., 2021), language spoken at home different from the language of instruction (e.g., Gambi and De Witte, 2022).

# **5** Remedial policies

As shown in Table 6, a number of different initiatives have been adopted in several EU Member States in order to help students catch up from the learning losses suffered due to Covid-19. These include: tutoring programmes, summer schools, measures designed to strengthen parental engagement, extension of school time, etc. While several studies (e.g., Kortekaas-Rijlaarsdam et al., 2022) carry out a literature review in an attempt to investigate the effectiveness of such provisions based on pre-Covid-19 experiences, very little research is so far available on the success of policies that have actually been implemented during the pandemic. However, this limited evidence would seem to suggest that online tutoring (individual or in small groups) is especially beneficial to students from less advantaged backgrounds. Arriola et al. (2021) evaluate "Menttores", an innovative programme providing free, online, two-to-one afterschool tutoring for less advantaged children who have been hardest hit by the pandemic in Spain. Menttores was addressed to pupils aged 12 to 15 and lasted 8 weeks, with three 50-minute sessions a week. The majority of tutors (i.e., about 87%) were qualified secondary school teachers. Using a robust scientific methodology like randomized controlled trial (RCT), it is found that this tutoring programme had a positive impact on student academic performance and improved pupils' socio-emotional well-being and aspirations. Similar evidence about the effectiveness of online tutoring for more vulnerable students comes from Italy. Carlana and La Ferrara (2021) assess the impact of an intervention called "TOP" that provided free individual online tutoring to economically disadvantaged middle-school Italian students during the lockdown. Tutors were university students who volunteered for 3 to 6 hours per week between mid-April and the beginning of June 2020. Again, employing RCT, it is concluded that TOP was successful in boosting students' achievement and substantially improved their socio-emotional skills, aspirations, and psychological well-being.

The important role played by tutoring programmes during Covid-19 is also highlighted by Pérez-Jorge et al. (2020). They find that a group of students from the Faculty of Education of the University La Laguna (Spain) were satisfied with the tutoring models that were adopted in the academic year 2019-20, especially in light of the university's closure due to the pandemic.

Country	Mitigation policies
Belgium	138 summer schools in Flanders in the summer of 2020
	190 summer schools in Flanders in the summer of 2021
	10 million euro allocated to summer schools in Flanders in the summer of 2022 (The Brussels
	Time, 2022)
Austria	Corona support package for students-
	about 3 million additional support hours in the summer and in the winter semester by
	February 2022 (European Commission, 2022)
Netherlands	Tutoring, extended school time, individual support, homework assistance, exam training,
	parental engagement European Commission / EACEA / Eurydice, 2022)
Lithuania	Supplementary tutoring between September and December 2021 (Švietimo, Mokslo ir
	Sporto Ministerija, 2021)
	Additional teaching assistants in general education for children with SEN
Slovakia	Summer schools and institution of a special fund in 2021 to cover schools' pandemic-related
	expenses (European Commission / EACEA / Eurydice, 2022)
France	Organization of the Open School summer 2020 programme to increase the teaching time
	devoted to practical experience (CEDEFOP, 2020)
Italy	Extracurricular small group tutoring for students at risk of school failure (European
	Commission / EACEA / Eurydice, 2022)
Poland	Rapid-response teams comprising counsellors, school psychologists, tutors, social workers
	were created by the Ministry of Education and Science <sup>10</sup> to address mental health problems
	faced by students with learning difficulties following Covid-19

Table 6: Examples of Covid-19 educational remedial initiatives in selected EU Member States

<sup>&</sup>lt;sup>10</sup> <u>Psychological and pedagogical support programme for pupils and teachers in a pandemic - Ministry of Education and Science - Portal Gov.pl</u> (www.gov.pl)

Portugal	Launch of plan 21 23 Escola+ whose duration is two years and has a budget of 900 million euro (European Commission, 2022a)
Germany	Planning of a comprehensive tutoring program (Freundl, 2021)
Czechia	A large-scale national tutoring plan was launched in September 2021 (European
	Commission, 2022b)
Luxembourg	Summer school in the summer of 2020 (Pfeiffer, 2020)

However, in spite of the aforementioned remedial policies, results presented in Table 4 indicate that overall, not only have students not recovered from the learning losses, but they have fallen further behind. Four concurrent explanations can be put forward.

First, students might have experienced problems in re-engaging with education activities following the re-opening of schools. There is evidence showing that, after several months of remote schooling, students in the US have become more passive ad feel disengaged from their learning (Toth, 2021). The stress and anxiety stemming from the pandemic are likely to have caused a fall in student motivation and morale. Wang and Liu (2022) perform a meta-analysis looking at the impact of Covid-19 on anxiety symptoms among Chinese university students. They find that anxiety prevalence was higher in later stages of the pandemic compared to early stages (29.1% vs 17.2%). Additionally, the uncertainty of the learning environment under Covid-19 could have also contributed to reduce students' educational aspirations (OECD, 2020a).

Second, as briefly mentioned above, the pandemic is likely to have exacerbated teacher shortage, exerting a negative effect on student achievement. While an increasing number of teachers have left the profession due to increased stress and workload as a result of Covid-19 (Pressley, 2021), there are not enough new teachers replacing them. Covid-19 could have contributed to increase the lack of attractiveness of the teacher profession. Teacher shortage is found to be correlated with poorer student performance in math, reading and science (OECD, 2020b).

Third, evidence from the US shows that planned tutoring programmes did not deliver what they originally intended in terms of student participation as well as in terms of number and duration of sessions. For instance, Carbonari et al. (2022) look at the implementation of tutoring programmes in 12 school districts across 10 states during the school year 2021-22. They find that in several districts the length and number of tutoring sessions turned out to be less than what was initially scheduled.

Fourth, it is possible that those students who most need remedial education are not taking advantage of existing catch-up programmes. Postlbauer et al. (2022) show that in Austria the intention to use remedial measures is strongly affected by family socio-economic background. Parents with a high level of education and a higher social status tend to show more interest for their children to participate in remedial courses.

Finally, evidence from the US suggests that most parents do not seem to understand the long-lasting consequences of the educational impact of Covid-19. A survey conducted in the autumn of 2022<sup>11</sup> shows that only about a quarter of the interviewed parents believe that their children's education is still behind due to the pandemic. This result is consistent with the findings of a similar survey carried out in March 2023<sup>12</sup>. According to this survey, more than 90% of parents think that their child has already or would soon catch up.

<sup>&</sup>lt;sup>11</sup> https://www.pewresearch.org/wp-content/uploads/2022/10/W115-topline\_COVID-impact-on-K-12-new.pdf

<sup>&</sup>lt;sup>12</sup> https://bealearninghero.org/wp-content/uploads/2023/04/Parents-2023-Go-Beyond-Grades.pdf

# 6 Conclusions

While there are already a large number of studies on the impact of Covid-19 on student performance, this is a key policy area that requires constant attention and monitoring. It is important to be aware that pupils will not automatically bounce back from missed learning following school closure during Covid-19. The negative effects of the pandemic on student learning are likely to persist without effective remedial policy action and they can have a long-term adverse impact on students and the economy<sup>13</sup>.

The meta-analysis carried out in this paper, which combines and summaries results from 55 studies covering 21 OECD countries, supports the conclusion that learning losses risk becoming persistent and are hard to make up. The results of our moderator analyses indicate that students have fallen further behind during the pandemic. On average, learning losses are found to be greater in assessments that occurred several months or more than one year after the first lockdown compared to assessments that took place in the early stages of the pandemic. This outcome contrasts with results found in previous meta-analyses (Di Pietro, 2023; Betthäuser et al., 2023) concluding that students did not lose any additional ground but failed to rebound.

Several explanations can be advanced to explain why pandemic learning losses are getting worse. It might have been hard for many students to go back to in-person classes after the lockdown. Some time may be needed for them to re-adjust to school life and re-engage with face-to-face learning. Covid-19 has resulted in increased teacher shortages, thereby negatively affecting student performance. Evidence from the US suggests also that it is quite possible that some remedial programmes are not producing the intended outcomes. This may be because the intensity and frequency of the remedial actions are not as high as planned or because these initiatives had a lower participation than expected. There are also some signs that a lot of parents in the US seem to underestimate the seriousness of the educational damage created by Covid-19.

However, one word of caution is needed here. Our results do not suggest that investments in education following Covid-19 are ineffective. It takes time for investments in education to deliver results, and most of those assessments which in our analysis are considered to have taken place in later stages of the pandemic have occurred shortly after or even before the adoption of remedial actions. This means that our estimates are unlikely to capture the full effect of these provisions. Such a consideration stresses the importance of closely monitoring changes in learning losses over time by constantly including in the analysis new studies that are more likely to analyse results from more recent assessments. In the spirit of the European Commission's "Learning Lab on Investing in Quality Education and Training"<sup>14</sup> initiative, frequently updated findings may contribute to designing appropriate policies, programmes and reforms which can make EU educational systems more effective, efficient and equitable (European Commission, 2023).

The results of our meta-analysis indicate also that students in OECD EU countries suffered smaller learning deficits than their peers in OECD non-EU countries. This result appears to be somehow in line with that obtained by Di Pietro (2023) who finds that the pandemic made European students lose less ground than their peers in non-European countries, although in his analysis he also considers students from several less developed countries in Africa and South America. The shorter duration of exposure to physical school closure for students in OECD EU countries relative to those in OECD non-EU countries could contribute to explaining why Covid-19 has had a smaller impact on the learning of the former group compared to the latter.

Finally, while our results confirm the findings of previous studies (including meta-analyses) that Covid-19 had, on average, an adverse effect on learning (between -0.11 and -0.17 standard deviations of student achievement), pupils from less advantaged backgrounds were hit especially hard. On the other hand, the collected evidence seems to suggest that are no statistically significant differences in the way the pandemic impacted the achievement of male and female students.

<sup>&</sup>lt;sup>13</sup> A report published by the World Bank, UNESCO and UNICEF (2021) states that, if learning losses are not recovered, there is the risk that this generation of students could lose 17 trillion dollars in lifetime earnings in present value, which is equivalent to approximately 14% of today's global GDP.

<sup>&</sup>lt;sup>14</sup> https://education.ec.europa.eu/focus-topics/improving-quality/learning-lab

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# List of abbreviations and definitions

CEDFOP	European Centre for the Development of Vocational Training
DL	DerSimonian and Laird
EACEA	European Education and Culture Executive Agency
FE	Fixed Effects
OECD	Organisation for Economic Cooperation and Development
PM	Paule-Mandel
RCT	Randomized Control Trial
RE	Random Effects
RML	Restricted Maximum Likelihood
RVE	Robust Variance Estimation
SJ	Sidik-Jonkman
ТОР	Tutoring Online Programme
UNESCO	United Nations Educational, Scientific and Cultural Organisation

UNICEF United Nations Children's Fund

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